

Perimeter Variance of Uniform Random Triangles

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ABSTRACT. Let T be a random triangle in a disk D of radius R (meaning that vertices are independent and uniform in D). We determine the bivariate density for two arbitrary sides a, b of T . In particular, we compute that $E(ab) = (0.837\dots)R^2$, which implies that $\text{Var}(\text{perimeter}) = (0.649\dots)R^2$. No closed-form expression for either coefficient is known. The Catalan numbers also arise here.

Let A, B, C denote three independent uniformly distributed points in the disk $D = \{(\xi, \eta) : \xi^2 + \eta^2 \leq R^2\}$. Let T denote the triangle with sides a, b, c opposite the vertices A, B, C . We are interested in the perimeter $a + b + c$ of triangle T . The univariate density $f(x)$ for side a is [1, 2, 3, 4, 5, 6, 7]

$$\frac{4x}{\pi R^2} \arccos\left(\frac{x}{2R}\right) - \frac{x^2}{\pi R^4} \sqrt{4R^2 - x^2}, \quad 0 < x < 2R$$

and

$$E(a) = \frac{128}{45\pi}R = (0.9054147873672267990407609\dots)R, \quad E(a^2) = R^2.$$

Clearly

$$E(\text{perimeter}) = 3E(a) = \frac{128}{15\pi}R = (2.7162443621016803971222828\dots)R$$

but to compute $\text{Var}(\text{perimeter}) = E(\text{perimeter}^2) - E(\text{perimeter})^2$, we will further need to consider cross-correlation ρ between sides.

The bivariate density $f(x, y)$ for sides a, b is

$$f(x, y) = \begin{cases} \varphi(x, y) & \text{if } x + y \leq 2R, \\ \psi(x, y) & \text{if } x + y > 2R \text{ and } x \leq 2R \end{cases}$$

when $0 \leq y \leq x$ (use symmetry otherwise) where

$$\begin{aligned} \varphi(x, y) = & \frac{2xy}{\pi R^6} \left\{ -\sqrt{(2R - x - y)(x - y)(2R + x - y)(x + y)} + \right. \\ & \left. 2(R - y)^2 \arccos\left(\frac{x^2 - 2Ry + y^2}{2x(R - y)}\right) + 2R^2 \arccos\left(\frac{x^2 + 2Ry - y^2}{2Rx}\right) \right\} + \\ & \frac{8xy}{\pi^2 R^6} \int_{R-y}^R t \arccos\left(\frac{t^2 + x^2 - R^2}{2tx}\right) \arccos\left(\frac{t^2 + y^2 - R^2}{2ty}\right) dt, \end{aligned}$$

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$$\psi(x, y) = \frac{8xy}{\pi^2 R^6} \int_{x-R}^R t \arccos\left(\frac{t^2 + x^2 - R^2}{2tx}\right) \arccos\left(\frac{t^2 + y^2 - R^2}{2ty}\right) dt.$$

It follows by numerical integration that

$$E(ab) = (0.8378520652962219016710654\dots)R^2$$

hence

$$\rho(a, b) = \frac{E(ab) - E(a)E(b)}{\sqrt{\text{Var}(a)\text{Var}(b)}} = 0.1002980835659001715822627\dots,$$

$$E(\text{perimeter}^2) = 3E(a^2) + 6E(ab) = (8.0271123917773314100263929\dots)R^2,$$

$$\text{Var}(\text{perimeter}) = (0.6491289571281667551974101\dots)R^2.$$

Exact evaluation of $E(ab)$ remains an open problem. We review derivation of the univariate case in section 1, imitating the analysis in [8, 9] very closely. (Parry's thesis [8] is concerned with triangles in three-dimensional space; it is surprising that our two-dimensional analog has not yet been examined.) The bivariate case is covered in section 2. An experimental consequence of our work is the formula

$$E(a^2 b^2) = \frac{13}{12}R^4$$

which we prove via a different approach in section 3. Finally, in section 4, the Catalan numbers from combinatorics appear rather unexpectedly.

1. UNIVARIATE CASE

We omit geometric details, referring to [8, 9] instead. The distance t between point C and the origin has density $2t/R^2$ for $0 < t < R$. Let $f(x|t)$ be the conditional density for distance x between points C and B , given t . We will compute the sought-after density $f(x)$ for side a via

$$f(x) = \int_0^R f(x|t) \frac{2t}{R^2} dt.$$

There are two subcases.

1.1. $0 < x < R$.

$$f(x) = \int_0^{R-x} \frac{2x}{R^2} \frac{2t}{R^2} dt + \int_{R-x}^R \frac{2x}{\pi R^2} \arccos\left(\frac{t^2 + x^2 - R^2}{2tx}\right) \frac{2t}{R^2} dt$$

which corresponds to formula (1.12) in Parry's thesis [8]. The \arccos term arises since, if the portion of a circle of radius x , center C contained within D has arclength $2\theta x$, then $f(x|t) = (2\theta x)/(\pi R^2)$; the Law of Cosines gives θ .

1.2. $R < x < 2R$.

$$f(x) = \int_{x-R}^R \frac{2x}{\pi R^2} \arccos\left(\frac{t^2 + x^2 - R^2}{2tx}\right) \frac{2t}{R^2} dt$$

which corresponds to Parry's (1.18). Straightforward integration provides the desired result (valid in both of the preceding regions).

2. BIVARIATE CASE

We omit geometric details, referring to [8] instead. The distance t between point C and the origin has density $2t/R^2$ for $0 < t < R$. Let $f(x, y | t)$ be the conditional density for distance x between points B and C , and distance y between points A and C , given t . We will compute the sought-after density $f(x, y)$ for sides a, b via

$$f(x, y) = \int_0^R f(x, y | t) \frac{2t}{R^2} dt.$$

There are six subcases.

2.1. $y < x$ and $0 < x < R$.

$$\begin{aligned} f(x, y) = & \int_0^{R-x} \frac{2x}{R^2} \frac{2y}{R^2} \frac{2t}{R^2} dt + \int_{R-x}^{R-y} \frac{2x}{\pi R^2} \arccos\left(\frac{t^2 + x^2 - R^2}{2tx}\right) \frac{2y}{R^2} \frac{2t}{R^2} dt + \\ & \int_{R-y}^R \frac{2x}{\pi R^2} \arccos\left(\frac{t^2 + x^2 - R^2}{2tx}\right) \frac{2y}{\pi R^2} \arccos\left(\frac{t^2 + y^2 - R^2}{2ty}\right) \frac{2t}{R^2} dt \end{aligned}$$

which corresponds to formula (4.26) in Parry's thesis [8].

2.2. $R < x < 2R$ and $0 < y < 2R - x$.

$$\begin{aligned} f(x, y) = & \int_{x-R}^{R-y} \frac{2x}{\pi R^2} \arccos\left(\frac{t^2 + x^2 - R^2}{2tx}\right) \frac{2y}{R^2} \frac{2t}{R^2} dt + \\ & \int_{R-y}^R \frac{2x}{\pi R^2} \arccos\left(\frac{t^2 + x^2 - R^2}{2tx}\right) \frac{2y}{\pi R^2} \arccos\left(\frac{t^2 + y^2 - R^2}{2ty}\right) \frac{2t}{R^2} dt \end{aligned}$$

which corresponds to Parry's (4.29). Straightforward integration gives $\varphi(x, y)$ (valid in both of the preceding regions).

2.3. $R < x < 2R$ and $2R-x < y < x$.

$$f(x, y) = \int_{x-R}^R \frac{2x}{\pi R^2} \arccos\left(\frac{t^2 + x^2 - R^2}{2tx}\right) \frac{2y}{\pi R^2} \arccos\left(\frac{t^2 + y^2 - R^2}{2ty}\right) \frac{2t}{R^2} dt$$

which corresponds to Parry's (4.32). This is, of course, $\psi(x, y)$.

2.4. $x < y$ and $0 < y < R$.

$$\begin{aligned} f(x, y) = & \int_0^{R-y} \frac{2x}{R^2} \frac{2y}{R^2} \frac{2t}{R^2} dt + \int_{R-y}^{R-x} \frac{2x}{R^2} \frac{2y}{\pi R^2} \arccos\left(\frac{t^2 + y^2 - R^2}{2ty}\right) \frac{2t}{R^2} dt + \\ & \int_{R-x}^R \frac{2x}{\pi R^2} \arccos\left(\frac{t^2 + x^2 - R^2}{2tx}\right) \frac{2y}{\pi R^2} \arccos\left(\frac{t^2 + y^2 - R^2}{2ty}\right) \frac{2t}{R^2} dt \end{aligned}$$

which corresponds to Parry's (4.35). This is, of course, $\varphi(y, x)$.

2.5. $R < y < 2R$ and $0 < x < 2R - y$.

$$\begin{aligned} f(x, y) = & \int_{y-R}^{R-x} \frac{2x}{R^2} \frac{2y}{\pi R^2} \arccos\left(\frac{t^2 + y^2 - R^2}{2ty}\right) \frac{2t}{R^2} dt + \\ & \int_{R-x}^R \frac{2x}{\pi R^2} \arccos\left(\frac{t^2 + x^2 - R^2}{2tx}\right) \frac{2y}{\pi R^2} \arccos\left(\frac{t^2 + y^2 - R^2}{2ty}\right) \frac{2t}{R^2} dt \end{aligned}$$

which corresponds to Parry's (4.38). This is, of course, $\varphi(y, x)$.

2.6. $R < y < 2R$ and $2R-y < x < y$.

$$f(x, y) = \int_{y-R}^R \frac{2x}{\pi R^2} \arccos\left(\frac{t^2 + x^2 - R^2}{2tx}\right) \frac{2y}{\pi R^2} \arccos\left(\frac{t^2 + y^2 - R^2}{2ty}\right) \frac{2t}{R^2} dt$$

which corresponds to Parry's (4.41). This is, of course, $\psi(y, x)$.

3. CHARACTERISTIC FUNCTION

We follow an approach found in [10, 11]. Let u, v, w denote the squared distances between A, B, C and the origin O . Let φ denote the angle between vectors \vec{OA}, \vec{OB} and ψ denote the angle between vectors \vec{OA}, \vec{OC} . We have

$$a^2 = v + w - 2\sqrt{vw} \cos(\psi - \varphi),$$

$$\begin{aligned} b^2 &= u + w - 2\sqrt{uw} \cos(\psi), \\ c^2 &= u + v - 2\sqrt{uv} \cos(\varphi) \end{aligned}$$

by the Law of Cosines, where u, v, w are independent uniform on $[0, R^2]$ and φ, ψ are independent uniform on $[0, 2\pi]$. The characteristic function for (a^2, b^2, c^2) is thus

$$\begin{aligned} g(r, s, t) &= \frac{1}{R^6} \frac{1}{4\pi^2} \int_0^{2\pi} \int_0^{2\pi} d\varphi d\psi \int_0^{R^2} \int_0^{R^2} \int_0^{R^2} du dv dw \exp \left[ir \left(v + w - 2\sqrt{vw} \cos(\psi - \varphi) \right) + \right. \\ &\quad \left. is \left(u + w - 2\sqrt{uw} \cos(\psi) \right) + it \left(u + v - 2\sqrt{uv} \cos(\varphi) \right) \right]. \end{aligned}$$

It is well-known that

$$E(c^2) = \frac{1}{i} \frac{\partial g}{\partial t} \Big|_{r=s=t=0}, \quad E(b^2 c^2) = \frac{1}{i^2} \frac{\partial^2 g}{\partial s \partial t} \Big|_{r=s=t=0}$$

and the former becomes

$$\begin{aligned} E(c^2) &= \frac{1}{i} \frac{\partial}{\partial t} \frac{1}{R^4} \frac{1}{2\pi} \int_0^{2\pi} d\varphi \int_0^{R^2} \int_0^{R^2} du dv \exp \left[it \left(u + v - 2\sqrt{uv} \cos(\varphi) \right) \right] \Big|_{t=0} \\ &= \frac{1}{i} \frac{\partial}{\partial t} \frac{1}{R^4} \int_0^{R^2} \int_0^{R^2} \exp(it(u+v)) J_0(2t\sqrt{uv}) du dv \Big|_{t=0} \end{aligned}$$

where $J_0(\theta)$ is the zeroth Bessel function of the first kind; hence

$$\begin{aligned} E(c^2) &= \frac{1}{i} \frac{1}{R^4} \int_0^{R^2} \int_0^{R^2} \frac{\partial}{\partial t} \exp(it(u+v)) J_0(2t\sqrt{uv}) \Big|_{t=0} du dv \\ &= \frac{1}{i} \frac{1}{R^4} \int_0^{R^2} \int_0^{R^2} i(u+v) du dv = R^2 \end{aligned}$$

which is consistent with before. The latter becomes

$$\begin{aligned} E(b^2 c^2) &= \frac{1}{i^2} \frac{\partial^2}{\partial s \partial t} \frac{1}{R^6} \frac{1}{4\pi^2} \int_0^{2\pi} \int_0^{2\pi} d\varphi d\psi \int_0^{R^2} \int_0^{R^2} \int_0^{R^2} du dv dw \exp \left[is \left(u + w - 2\sqrt{uw} \cos(\psi) \right) + \right. \\ &\quad \left. it \left(u + v - 2\sqrt{uv} \cos(\varphi) \right) \right] \Big|_{s=t=0} \\ &= - \frac{\partial^2}{\partial s \partial t} \frac{1}{R^6} \int_0^{R^2} \int_0^{R^2} \int_0^{R^2} \exp(is(u+w)) J_0(2s\sqrt{uw}) \exp(it(u+v)) J_0(2t\sqrt{uv}) du dv dw \Big|_{s=t=0} \end{aligned}$$

$$\begin{aligned}
&= -\frac{1}{R^6} \int_0^{R^2} \int_0^{R^2} \int_0^{R^2} \frac{\partial^2}{\partial s \partial t} \exp(is(u+w)) J_0(2s\sqrt{uw}) \exp(it(u+v)) J_0(2t\sqrt{uv}) \Big|_{s=t=0} du dv dw \\
&= -\frac{1}{R^6} \int_0^{R^2} \int_0^{R^2} \int_0^{R^2} -(u+v)(u+w) du dv dw = \frac{13}{12} R^4
\end{aligned}$$

as was to be shown. The fact that $13/12 - 1 = 1/12 \neq 0$ offers the simplest proof we know that arbitrary sides of a random triangle in D must be dependent.

4. CATALAN NUMBERS

Let $R = 1$ for the remainder of our discussion. From

$$P(a^2 < x) = P(a < \sqrt{x}) = \int_0^{\sqrt{x}} \left(\frac{4\xi}{\pi} \arccos\left(\frac{\xi}{2}\right) - \frac{\xi^2}{\pi} \sqrt{4-\xi^2} \right) d\xi$$

we obtain that the density for a^2 is

$$\left(\frac{4\sqrt{x}}{\pi} \arccos\left(\frac{\sqrt{x}}{2}\right) - \frac{x}{\pi} \sqrt{4-x} \right) \frac{1}{2\sqrt{x}}, \quad 0 < x < 4.$$

On the one hand, the characteristic function for a^2 is

$$\int_0^1 \int_0^1 \exp(it(u+v)) J_0(2t\sqrt{uv}) du dv$$

by the preceding section; on the other hand, it is

$$\begin{aligned}
&\int_0^4 \exp(itx) \left(\frac{4\sqrt{x}}{\pi} \arccos\left(\frac{\sqrt{x}}{2}\right) - \frac{x}{\pi} \sqrt{4-x} \right) \frac{1}{2\sqrt{x}} dx \\
&= \frac{i}{t} [1 - \exp(2it) (J_0(2t) - iJ_1(2t))] \\
&= \frac{i}{t} [1 - h(t)]
\end{aligned}$$

where $J_1(\theta) = -J'_0(\theta)$. A direct evaluation of the double integral seems to be difficult. Boersma [12], using work of Zernike & Nijboer [13, 14, 15], gave a rapidly-convergent series for the inner integral:

$$\int_0^1 \exp(itu) J_0(2t\sqrt{uv}) du = \frac{\sqrt{\pi}}{t^{3/2}v^{1/2}} \exp\left(\frac{it}{2}\right) \sum_{n=0}^{\infty} (-i)^n (2n+1) J_{n+1/2}\left(\frac{t}{2}\right) J_{2n+1}(2t\sqrt{v})$$

but this apparently does not help with the outer integral.

Let $I_0(\theta)$ be the zeroth modified Bessel function of the first kind and $I_1(\theta) = I_0'(\theta)$. We note that the exponential generating function for the Catalan numbers [16]:

$$\exp(2t) (I_0(2t) - I_1(2t)) = \sum_{n=0}^{\infty} \frac{1}{(n+1)!} \binom{2n}{n} t^n$$

is remarkably similar to the expression for $h(t)$. Replacing t by it , we obtain

$$h(t) = \exp(2it) (J_0(2t) - iJ_1(2t)) = \sum_{n=0}^{\infty} \frac{1}{(n+1)!} \binom{2n}{n} (it)^n$$

because $J_0(i\theta) = I_0(\theta)$, $J_1(i\theta) = iI_1(\theta)$. Therefore the Catalan numbers are associated with the characteristic function for a^2 . We wonder if a two-dimensional integer array, suitably generalizing the Catalan numbers, can be associated with the characteristic function for (a^2, b^2) :

$$\int_0^1 \int_0^1 \int_0^1 \exp(is(u+w)) J_0(2s\sqrt{uw}) \exp(it(u+v)) J_0(2t\sqrt{uv}) du dv dw.$$

Since the bivariate density $f(x, y)$ for (a, b) is much more complicated than the univariate density $f(x)$ for a , an answer to our question may be a long time coming.

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REFERENCES

- [1] R. Deltheil, *Probabilités Géométriques*, t. 2, *Traité du calcul des Probabilités et de ses Applications*, f. 2, ed E. Borel, Gauthier-Villars, 1926, pp. 40–42, 114–120.
- [2] J. M. Hammersley, The distribution of distance in a hypersphere, *Annals Math. Statist.* 21 (1950) 447–452; MR0037481 (12,268e).
- [3] R. D. Lord, The distribution of distance in a hypersphere, *Annals Math. Statist.* 25 (1954) 794–798; MR0065048 (16,377d).
- [4] V. S. Alagar, The distribution of the distance between random points, *J. Appl. Probab.* 13 (1976) 558–566; MR0418183 (54 2#6225).
- [5] H. Solomon, *Geometric Probability*, SIAM, 1978, pp. 35–36, 128–129; MR0488215 (58 #7777).

- [6] S. R. Dunbar, The average distance between points in geometric figures, *College Math. J.* 28 (1997) 187–197; MR1444006 (98a:52007).
- [7] S.-J. Tu and E. Fischbach, Random distance distribution for spherical objects: general theory and applications to physics, *J. Phys. A* 35 (2002) 6557–6570; MR1928848.
- [8] M. Parry, *Application of Geometric Probability Techniques to Elementary Particle and Nuclear Physics*, Ph.D. thesis, Purdue Univ., 1998.
- [9] M. Parry and E. Fischbach, Probability distribution of distance in a uniform ellipsoid: theory and applications to physics, *J. Math. Phys.* 41 (2000) 2417–2433; MR1751899 (2001j:81267).
- [10] Y. Isokawa, Limit distributions of random triangles in hyperbolic planes, *Bull. Faculty Educ. Kagoshima Univ. Natur. Sci.* 49 (1997) 1–16; MR1653095 (99k:60020).
- [11] Y. Isokawa, Geometric probabilities concerning large random triangles in the hyperbolic plane, *Kodai Math. J.* 23 (2000) 171–186; MR1768179 (2001f:60014).
- [12] J. Boersma, On the computation of Lommel’s functions of two variables, *Math. Comp.* 16 (1962) 232–238; MR0146419 (26 #3941).
- [13] F. Zernike and B. R. A. Nijboer, Théorie de la diffraction des aberrations, *La Théorie des Images Optiques*, Proc. 1946 Paris colloq., ed. P. Fleury, A. Maréchal and C. Anglade, La Revue d’Optique, 1949, pp. 227–235.
- [14] B. R. A. Nijboer, *The Diffraction Theory of Aberrations*, Ph.D. thesis, Univ. of Groningen, 1942, available online at http://www.nijboerzernike.nl/_html/intro.html.
- [15] A. J. E. M. Janssen, J. J. M. Braat and P. Dirksen, On the computation of the Nijboer-Zernike aberration integrals at arbitrary defocus, *J. Mod. Optics* 51 (2004) 687–703; available online at http://www.nijboerzernike.nl/_html/biblio.html.
- [16] N. J. A. Sloane, On-Line Encyclopedia of Integer Sequences, A144186.
- [17] S. Finch, Random triangles. I–IV, unpublished essays (2010), <http://algo.inria.fr/bsolve/>.

- [18] S. Finch, Simulations in R involving triangles and tetrahedra,
<http://algo.inria.fr/csolve/rsimul.html>.

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